# **Better Bunching, Nicer Notching**

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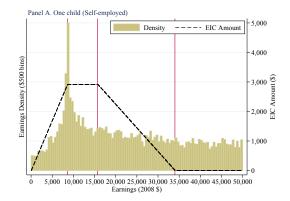
EEA-ESEM, Manchester

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### Goal: Estimate the Tax Elasticity of Reported Income.

Saez [2010] and Kleven and Waseem [2013]'s insight is that the elasticity relates to amount of bunchers.

More bunchers = larger elasticity; *ceteris paribus*.



## The Bunching Estimator

- Ideally, you would like to observe the same distribution of individuals facing different tax changes
  - There are not that many quasi experiments.
- The bunching estimator is an appealing identification strategy because it only takes one population of individuals and one (piece-wise linear) budget set to estimate the elasticity.
- Many examples of applications with piece-wise linear budget sets: Prescription drug insurance [Einav et al., 2017], pensions systems [Brown and Laschever, 2012], welfare programs [Camacho and Conover, 2011], education policy [Dee et al., 2011], labor regulations [Garicano et al., 2016], minimum wages [Dube et al., 2017], fuel economy policy [Sallee and Slemrod, 2012], real estate taxes [Kopczuk and Munroe, 2015], and pricing schemes in electricity [Ito, 2014], cellular service [Huang, 2008], and water markets [Olmstead et al., 2007].
- Our focus is on one population facing one budget set.
  - There are other estimators using variation from piece-wise linear budget constraints e.g., Blomquist and Newey [2002].

# What Do We Learn From Bunching?

- This paper clarifies that non-parametric identification of the elasticity is impossible in the case of kinks.
- We propose identifying conditions that are weaker than those used before.

Solution 1 (Bounds)

- non-parametric shape restriction: partial identification

Solution 2 (Truncated Tobit)

- covariates and semi-parametric distribution restriction: point identification

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weakest assumptions

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strongest assumptions

### **Utility Maximization Problem**

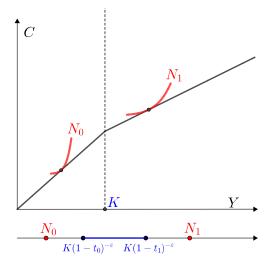
Agent type N\* maximizes utility U(C, Y; N\*) choosing consumption C and labor income Y

$$\max_{C,Y} \qquad C - \frac{N^*}{1 + 1/\varepsilon} \left(\frac{Y}{N^*}\right)^{1 + \frac{1}{\varepsilon}}$$

s.t. 
$$C = \mathbb{I}\{Y \le K\}[I_0 + (1 - t_0)Y] + \mathbb{I}\{Y > K\}[I_1 + (1 - t_1)(Y - K)]$$

- Piece-wise linear budget with intercept  $I_j$  and slope  $1 t_j$
- There is a tax rate change  $t_0 < t_1$  at Y = K (kink)

## Solution to Utility Maximization Problem



## Goal: Invert Solution Mapping to Get $\varepsilon$

The solution in logs is:

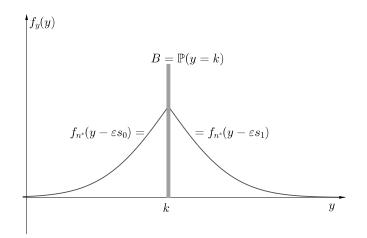
$$y = \begin{cases} n^* + \varepsilon s_0 & \text{, if } 0 < n^* < k - \varepsilon s_0 \\ k & \text{, if } k - \varepsilon s_0 \leq n^* \leq k - \varepsilon s_1 \\ n^* + \varepsilon s_1 & \text{, if } k - \varepsilon s_1 < n^* < \infty \end{cases}$$
(1)

where:

$$y = \log(Y)$$
,  $n^* = \log(N^*)$ ,  $k = \log(K)$ ,  $s_j = \log(1 - t_j)$ ,

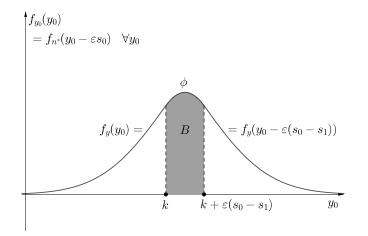
- For the terms of the counterfactual scenario of no tax change:  $t_0 = t_1$
- Counterfactual income :  $y_0 = n^* + \varepsilon s_0$

#### Distribution of Income Observed Data



## **Distribution of Income**

Unobserved Counterfactual



### How Does the Literature Identify the Elasticity?

- The 'bunching estimator' of Saez [2010] uses the definition of B

$$B = \int_{k}^{k+\varepsilon(s_{0}-s_{1})} f_{y_{0}}\left(y\right) \, dy,$$

- makes a trapezoidal approximation,

$$\cong 0.5 \left[ f_{y_0} \left( k + \varepsilon \left( s_0 - s_1 \right) \right) + f_{y_0} \left( k \right) \right] \varepsilon \left( s_0 - s_1 \right),$$

- replaces  $f_{y_0}$  by  $f_y$ ,

$$= 0.5 \left[ f_y(k^+) + f_y(k^-) \right] \varepsilon \left( s_0 - s_1 \right),$$

- and solves for the elasticity

$$\varepsilon \cong B / \left\{ (s_0 - s_1) \, 0.5 \left[ f_y(k^+) + f_y(k^-) \right] \right\}$$

- This restricts the PDF  $f_{n^*}$  to be "linear" for  $n^* \in [k - \varepsilon s_0, k - \varepsilon s_1]$ .

Literature Relies on Strong Functional Form Assumptions

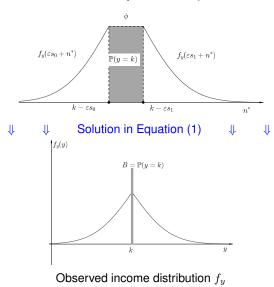
- "But isn't linearity approximately true if the bunching interval is small?"

- There is no way to know because the length of the interval depends on the  $\underline{\mathsf{unknown}} \ \varepsilon.$ 

- Chetty et al. [2011] make a stronger functional form assumption:  $n^*$  has a flat PDF inside the bunching interval.
- These are not non-parametric identification strategies.
- We find  $\varepsilon$  is very sensitive to the shape of  $f_{n^*}$ .

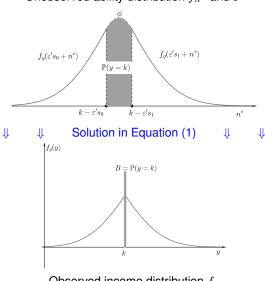
## Identification without Restrictions on $f_{n^*}$ is Impossible

Unobserved ability distribution  $f_{n*}$  and  $\varepsilon$ 



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Unobserved ability distribution  $f_{n^*}$  and  $\varepsilon'$ 



## Data Justifies Any Elasticity

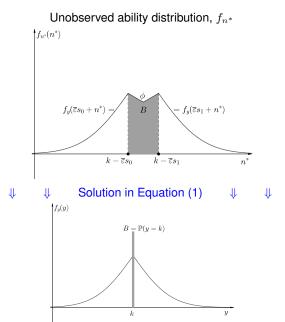
**Lemma:** Assume  $f_{n^*}$  is an unobserved continuous PDF with support  $(-\infty, +\infty)$ , and that we observe  $f_y$ .

Then, for every elasticity  $\varepsilon > 0$ , there exists a  $f_{n^*}$  such that Equation (1) maps the distribution of  $n^*$  into the distribution of y.

Therefore, it is impossible to identify the elasticity.

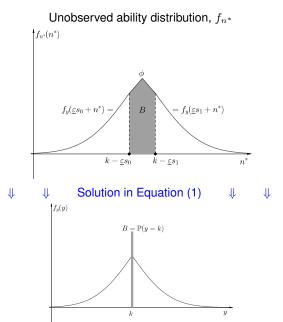
- There is an old literature on impossible inference.
  - For a review see Bertanha and Moreira (2019).
  - Here we have a worse problem: not even identification.
- This result also appears in Blomquist and Newey [2017].

#### Shape Restriction Yields Partial Identification Assume $f_{n^*}$ is Lipschitz with constant M



12

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### Partial Identification of $\varepsilon$

**Theorem:** Assume  $f_{n^*}$  is Lipschitz with known constant *M*.

Then,

$$\varepsilon \in \Upsilon = \left\{ \begin{array}{ll} \emptyset & \text{, if } B < \frac{\left|f_y(k^+) - f_y(k^-)\right| \left[f_y(k^+) + f_y(k^-)\right]}{2M} \\ \left[\underline{\varepsilon}, \overline{\varepsilon}\right] & \text{, if } \frac{\left|f_y(k^+) - f_y(k^-)\right| \left[f_y(k^+) + f_y(k^-)\right]}{2M} \le B < \frac{f_y(k^+)^2 + f_y(k^-)^2}{2M} \\ \left[\underline{\varepsilon}, \infty\right) & \text{, if } \frac{f_y(k^+)^2 + f_y(k^-)^2}{2M} \le B \end{array} \right. \right\}$$

where

$$\underline{\varepsilon} = \frac{2\left[f_y(k^+)^2/2 + f_y(k^-)^2/2 + M B\right]^{1/2} - \left(f_y(k^+) + f_y(k^-)\right)}{M(s_0 - s_1)}$$
$$\overline{\varepsilon} = \frac{-2\left[f_y(k^+)^2/2 + f_y(k^-)^2/2 - M B\right]^{1/2} + \left(f_y(k^+) + f_y(k^-)\right)}{M(s_0 - s_1)}$$

**Solution 1 (Bounds) :** We recommend a sensitivity analysis for various choices of M.

## Bunching is a Censored Regression Model

#### Classic Tobit

Suppose  $n^* \sim N(\mu, \sigma^2)$ .

Left censored at k:  $y = \max\{k; n^*\}$ , or

right censored at k:  $y = \min\{k; n^*\}$ .

Mid-censored Model: Equation (1) rewrites as

 $y = \min\{\varepsilon s_0 + n^*, \max\{k, \varepsilon s_1 + n^*\}\}.$ 

- The elasticity is the difference of intercepts divided by  $(s_1 - s_0)$ 

## Use Covariates to Predict Ability with a Tobit

#### Mid-censored Tobit

-There is a vector of covariates X

Idea: instead of polynomial or linear extrapolations, use the relationship between n\* and X for non-bunching people to predict the distribution of n\* for bunching people.

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Assumption: n^* \mid X \sim N(X\beta; \sigma^2)
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Question: what if conditional normality is incorrect?

- MLE is inconsistent ...

## Tobit is Robust to Non-normality

Assumption: there exists an unique  $(\beta, \sigma)$  such that  $F_{n^*}(n) = \mathbb{E}\left[\Phi\left(\frac{n-X\beta}{\sigma}\right)\right]$ , plus some regularity conditions.

- Semi-parametric class that does NOT require conditional normality.

**Lemma:** Let  $\theta^* = (\beta^*, \sigma^*, \varepsilon^*)$  be the probability limit of the Tobit MLE, and suppose the Assumption above holds.

If the true distribution  $F_y(y)$  matches the Tobit fitted distribution  $G_y(y; \theta^*)$ , then  $\varepsilon^*$  is equal to the true elasticity.

#### Solution 2 (Truncated Tobit):

- truncate the sample of y to a neighborhood of k;
- vary the size of the neighborhood, from full to smallest sample;
- examine how the fit and estimates vary.

# Application

Employ data originally studied by Saez [2010]

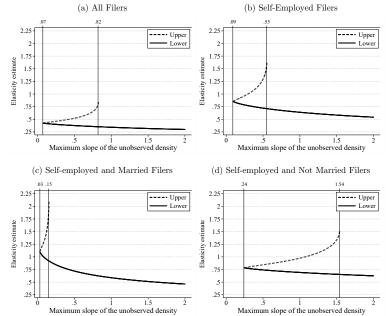
- Individual Public Use Tax Files constructed by the IRS
- Annual cross-sections 1995-2004

► Focus on \$8,580 kink in the EITC schedule:

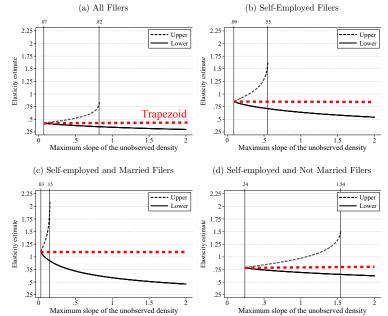
- tax rate changes from -34 % to 0%
- We constrast trapezoid-based estimates with
  - Solutions 1 (Bounds)
  - Solutions 2 (Truncated Tobit)

- dummy covariates: year effects, types of deductions, received social security benefits, used a tax prep software, etc.

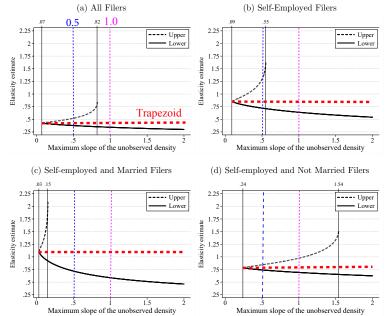
## Partial Identification



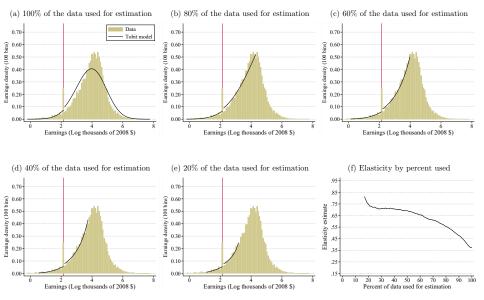
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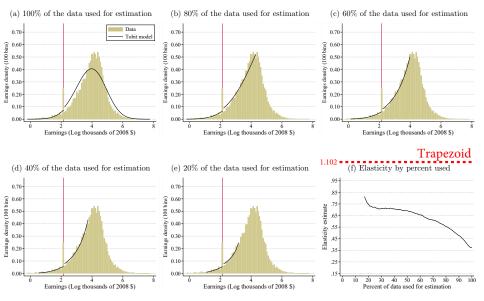
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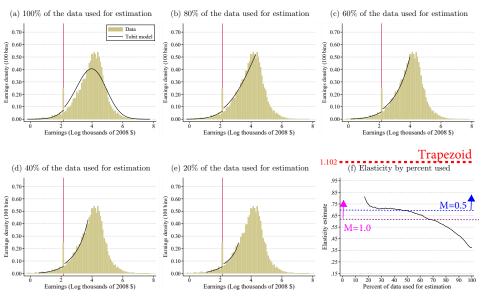
#### Truncated Tobit Self-employed Married



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#### Truncated Tobit Self-employed Married



# Conclusions

- Minimal restrictions partially identify the elasticity
- Connection between bunching and censored regressions allows for:
  - Covariates: more meaningful extrapolation
  - Semi-parametric restrictions on the distribution of  $n^*$

#### Future of Bunching:

- examine sensitivity of estimates using Bounds and Truncated Tobit
- STATA package coming soon!

#### More on the Paper, Skipped Today:

- failure of the widely used "polynomial strategy"
- multiple kinks and notches
- non-parametric identification is possible in the case of notches
- point indentification with censored quantile regressions